

Random Graphs

Foundations of Data Science, 8.1-8.2

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Note: I will write $x_n \rightarrow c$ for $x_n \xrightarrow{n \rightarrow \infty} c$ and $x_n \sim y_n$ for $\frac{x_n}{y_n} \rightarrow 1$.

1 Terminology

Definition 1. Let $n \in \mathbb{N}$, $p \in (0, 1)$. The Erdős-Rényi Graph $G(n, p)$ is a random graph where every possible edge is present with probability p (independent of other edges).

$G(n, p)$ is a random variable. We usually look at families of random graphs where $n \rightarrow \infty$ and p is a function of n . We say a property of such a family occurs *almost surely* if its probability converges to 1 as n approaches infinity.

2 General Observations

Even though the edges are chosen independently, certain global properties emerge for specific choices of p . Specifically, for some properties, there is a *threshold* p' , meaning if p is sufficiently smaller than p' , the property occurs with high (low) probability; and if p is sufficiently larger than p' , then the property occurs with low (high) probability. We will look at several properties with different thresholds.

3 Degree distribution

Example 1. In $G(n, 1/2)$, for $\varepsilon > 0$, the degree of each vertex is almost surely within $(1 \pm \varepsilon)n/2$.

Proof notes. The standard deviation is $\sigma = \sqrt{\text{Var}} = \frac{\sqrt{n-1}}{2}$. By Chebychev's inequality, we have $\Pr[|x - E[X]| \geq c\sigma] \leq 1/c^2$. Therefore, any constant fraction of the probability mass lies within $(n-1)/4 \pm c\sqrt{n-1}/2$, or within $(n-1)/4 \pm \varepsilon$ for large n . \square

For $G(n, p)$, the degree distribution is also a binomial distribution:

$$\begin{aligned} \Pr[\text{Vertex has degree } k] &= \binom{n-1}{k} p^k (1-p)^{n-k-1} \\ &= \frac{n-k}{n(1-p)} \binom{n}{k} p^k (1-p)^{n-k} \sim \binom{n}{k} p^k (1-p)^{n-k} \end{aligned}$$

for $k \in o(n)$ and $p \rightarrow 0$.

Vertex degrees in real-world graphs often have a power law distribution, where the probability of a degree k decreases polynomially in k , not exponentially as with a binomial distribution. We won't consider this here.

Theorem 1 (Theorem 8.1, Corollary 8.2 without approximations). *Let $\varepsilon \in (0, 1)$. If $p \geq 9 \frac{\ln n}{(n-1)\varepsilon^2}$, then almost surely every vertex has degree in the range $(1 \pm \varepsilon)(n-1)p$.*

Proof. The degree $\deg(v)$ is a sum of $m = n-1$ independent Bernoulli random variables x_1, x_2, \dots, x_m , where x_i is the indicator variable that the i^{th} edge from v is present. Applying the Chernov Bound (Theorem 12.6):

$$\Pr[|mp - \deg(v)| \geq \varepsilon mp] \leq 3e^{-mp\varepsilon^2/8} \leq 3n^{-9/8} \rightarrow 0.$$

\square

4 The binomial distribution approaches the poisson distribution

Lemma 1. *Assume $k \leq o(n)$ and $\binom{n}{k} \sim n^k/k!$. As n goes to infinity, the binomial distribution with d/n*

$$\Pr[k] = \binom{n}{k} \left(\frac{d}{n}\right)^k \left(1 - \frac{d}{n}\right)^{n-k}$$

approaches the poisson distribution

$$\Pr[k] = \frac{d^k}{k!} e^{-d}.$$

Therefore, in $G(n, d/n)$ for a constant d , the degree distribution is a poisson distribution.

Proof sketch. Note that $(1-d/n)^n \sim e^{-d}$. It remains to show that $(1-d/n)^k \rightarrow 1$. Clearly, $(1-d/n)^k < 1$. By Bernoulli's inequality, for $n \geq d$,

$$\left(1 - \frac{d}{n}\right) \geq 1 - \frac{kd}{n} \rightarrow 0.$$

□

The authors seem to claim that $\binom{n}{k} \sim n^k/k!$ follows from $k \in o(n)$, which is not clear to me. It certainly is true if k is a constant.

Example 2. In $G(n, 1/n)$, there is a vertex of degree $\Omega(\log n / \log \log n)$ with high (constant) probability if $n = k!$ for some k .

Proof notes. As argued in the book, $\log n / \log \log n \leq k$.

Let A_i be the event that the vertex i has degree k and let A be the event that any vertex has degree k . Then

$$\Pr[A] = \Pr\left[\bigcup_{i=0}^n A_i\right] \geq \sum_{i=1}^n \Pr[A_i] - \sum_{i < j} \Pr[A_i \cap A_j]$$

by the Bonferroni inequality¹. Let i, j be vertices, and B_{ij} the event that the edge $\{i, j\}$ exists. Then we can calculate

$$\begin{aligned} \Pr[A_i | B_{ij}] &\sim \frac{k}{en}, \\ \Pr[A_i | \overline{B_{ij}}] &\sim \frac{1}{en}, \text{ and therefore} \\ \Pr[A_i \cap A_j] &= \Pr[B_{ij}] \cdot \Pr[A_i \cap A_j | B_{ij}] + \Pr[\overline{B_{ij}}] \cdot \Pr[A_i \cap A_j | \overline{B_{ij}}] \\ &\sim \frac{1}{n} \cdot \left(\frac{k}{en}\right)^2 + \left(1 - \frac{1}{n}\right) \left(\frac{1}{en}\right)^2 \sim \frac{k^2}{e^2 n^3} + \frac{1}{e^2 n^2}, \end{aligned}$$

using independence of A_i and A_j under B_{ij} or $\overline{B_{ij}}$. Summing up, we get

$$\Pr[A] \gtrsim n \frac{1}{en} - n^2 \left(\frac{k^2}{e^2 n^3} + \frac{1}{e^2 n^2}\right) \sim \frac{1}{e} - \frac{1}{e^2}.$$

¹https://en.wikipedia.org/wiki/Boole%27s_inequality

5 Existence of Triangles

Lemma 2. *The expected number of triangles in $G(n, d/n)$ converges to $d^3/6$ from below.*

Proof notes. More precisely, if x_n is the number of triangles,

$$\mathbb{E}[x_n] = \binom{n}{3} p^3 = \binom{n}{3} \left(\frac{d^3}{n^3}\right) = \frac{d^3}{6} \cdot \frac{(n-1)(n-2)}{n^2}.$$

□

We now want to find the probability that $G(n, d/n)$ contains at least one triangle. From the lemma it not necessarily follows that this probability is high; a huge amount of triangles could be concentrated in a tiny number of graphs. But this would mean that the variance is high, which we will prove to be false.

Lemma 3. *Let x be the number of triangles in $G(n, d/n)$. Then $\text{Var}[x] \leq \mathbb{E}[x] + o(1)$.*

Proof notes. $\binom{n}{3} p^3$ is exactly the expectation value of x , therefore part 3 contributes $\mathbb{E}[x]$ to the whole sum and

$$\mathbb{E}[x^2] \leq \mathbb{E}^2[x] + \mathbb{E}[x] + o(1),$$

yielding the slightly stronger upper bound $\text{Var}[x] \leq \mathbb{E}[x] + o(1)$. □

Theorem 2. *Let x be the number of triangles in $G(n, d/n)$. Then*

$$\Pr[x = 0] \leq \frac{6}{d^3} + o(1).$$

Proof notes. By Chebychev's inequality,

$$\Pr[x = 0] \leq \frac{\text{Var}[x]}{\mathbb{E}^2[x]} \leq \mathbb{E}[x] + \frac{o(1)}{\mathbb{E}[x]} = \mathbb{E}[x] + o(1).$$

□

Thus, for $d > \sqrt[3]{6}$, $\Pr[x = 0] < 1$ and $G(n, d/n)$ has a triangle with nonzero probability, and for $p \in \omega(1/n)$, $G(n, d/n)$ almost surely has a triangle.² Furthermore, if $p \in o(1/n)$, $\mathbb{E}[x] \rightarrow 0$, so $G(n, p)$ almost surely has no triangle.

□

²Note that d is not actually required to be a constant in the above proofs, so $p \in \omega(1/n)$ can be written as $d \in \omega(1)$.

Table 1: Phase transitions

| Probability | Transition |
|-----------------------------------|--|
| $p < o(\frac{1}{n})$ | Forest of trees No component of size greater than $\mathcal{O}(\log n)$ |
| $p = \frac{d}{n}, d < 1$ | Cycles appear No component of size greater than $\mathcal{O}(\log n)$ |
| $p = \frac{1}{n}$ | Components of size $\mathcal{O}(n^{2/3})$ |
| $p = \frac{d}{n}, d > 1$ | Giant component plus $\mathcal{O}(\log n)$ components |
| $p = \frac{1}{2} \frac{\ln n}{n}$ | Giant component plus isolated vertices |
| $p = \frac{\ln n}{n}$ | Disappearance of isolated vertices Appearance of Hamilton circuit Diameter $\mathcal{O}(\log n)$ |
| $p = \sqrt{\frac{2 \ln n}{n}}$ | Diameter 2 |
| $p = \frac{1}{2}$ | Clique of size $(2 - \varepsilon) \ln n$ |

6 Phase Transitions

Besides the existence of triangles, there are other graph properties that arise more abruptly as p passes some threshold.

Definition 2. Given some property, if there exists a function $p(n)$ such that

- for all $p_1(n) \leq o(p(n))$, $G(n, p_1(n))$ almost surely does not have the property; and
- for all $p_2(n) \geq \omega(p(n))$, $G(n, p_2(n))$ almost surely has the property,

then we say that a *phase transition* occurs, and $p(n)$ is the *threshold*.

If for $c < 1$, $G(n, cp(n))$ almost surely does not have the property and for $c > 1$, $G(n, cp(n))$ almost surely has the property, then $p(n)$ is a *sharp threshold*.

From the discussion in section 5, we know that the existence of triangles has a threshold at $p = 1/n$. Note that for non-sharp thresholds, multiplicative constants do not matter, so we could write the threshold as $\Theta(1/n)$.

To estimate the probability of the existence of something, we often use a variable x that *counts* the number of occurrences of an item in a random graph and calculate expected value and variance of x . This lets us apply the following two theorems:

Theorem 3 (First moment method). *Let x_n be a sequence of nonnegative integer random variables. If $E[x_n]$ goes to zero as n goes to infinity, $Pr[x_n = 0]$ goes to one, i.e. $x_n = 0$ almost surely is true.*

Therefore, if for some p the expected value of the item count goes to zero, the probability of the existence of an item also goes to zero.

Theorem 4 (Second moment method, Theorem 8.3). *Let x_n be a sequence of random variables. Let there be an $\varepsilon > 0$ such that $E[x_n] > \varepsilon$ for all n , i.e. $\liminf(x_n) > 0$. If $\text{Var}[x_n] \leq o(E^2[x_n])$, then x_n is almost surely greater than zero.*

Proof notes: The theorem in the book only requires $E[x] > 0$, which is either incorrect or written in a weird way. The condition $E[x_n] > 0$ for all n isn't sufficient, as $E[x_n]$ could still converge to 0 and then $\text{Var}[x]/E^2[x]$ not necessarily converges to 0.

Note that $\text{Var}[x_n] \leq o(E[x_n])$ is equivalent to $\sigma[x_n] \leq o(E[x_n])$, we require the standard deviation to be significantly smaller than the expectation.

Corollary 1 (Corollary 8.4). *Let x_n be a sequence of random variables as above. If $E[x^2] \leq E^2[x](1 + o(1))$, then x is almost surely greater than zero.*

The second moment method can be used to show that for some p , the probability of the existence of an item goes to one. Note that computing the expected value is often much easier than computing the variance (respectively $E[x^2]$).

7 Graph diameter two

A random graph having diameter less or equal two is an example of a property with sharp threshold.

Theorem 5 (Theorem 8.5). *The property "G(n, p) has diameter two" has a sharp threshold at $p = \sqrt{\frac{2 \ln n}{n}}$.*

Proof. Note that $\binom{n}{2} \sim n^2/2$ and

$$(1 - c^2 \ln n/n)^n \sim e^{-c^2 \ln n}$$

□

8 Disappearance of Isolated Vertices

Theorem 6 (Theorem 8.6). *The property "G(n,p) has isolated vertices" has a sharp threshold at $\frac{\ln n}{n}$.*

Proof notes. With $p = c\frac{\ln n}{n}$, we get

$$E[x] = \left(1 - \frac{c \ln n}{n}\right)^{n-1} \sim \left(1 - \frac{c \ln n}{n}\right)^n, \text{ as } \left(1 - \frac{c \ln n}{n}\right) \rightarrow 1.$$

□

9 Hamilton circuits

The existence of a hamilton circuit is a property with a threshold that can't be directly found with the first and second moment methods. At the value of p at which the expected *number* of hamilton circuits starts to go to infinity, these hamilton circuits are highly concentrated in a small number of graphs, i.e. the variance is high. This illustrates the necessity of the second moment method.

Theorem 7. *The expected number of hamilton circuits in $G(n, d/n)$ is 0 for $d \leq e$ and ∞ for $d > e$.*

Proof notes. Using Stirling's approximation:

$$\begin{aligned} E[x] &= \frac{1}{2}(n-1)! \left(\frac{d}{n}\right)^n = \frac{1}{2n} n! \left(\frac{d}{n}\right)^n \\ &\sim \frac{1}{2n} \sqrt{2\pi n} \left(\frac{n}{e}\right)^n \left(\frac{d}{n}\right)^n = \sqrt{\frac{\pi}{2n}} \left(\frac{d}{e}\right)^n \end{aligned}$$

□

Still, the threshold for the *existence* of hamilton circuits can't be $p = e/n$, as such graphs are with high probability not even connected. The actual threshold is $\frac{1}{n} \ln n$ (not proved here).